

Collective Matrix Completion

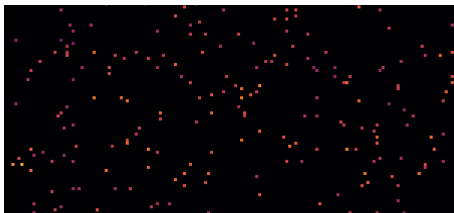
Mokhtar Z. Alaya

Modal'X, Paris Nanterre University

joint work with Olga Klopp (ESSEC Business School)



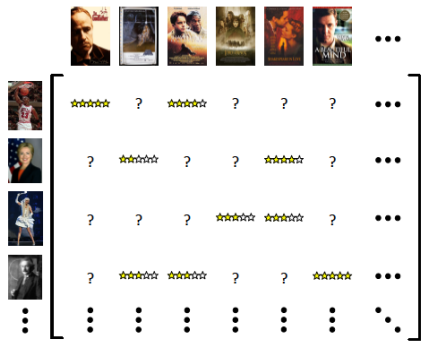
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- **Task:** given a partially observed data matrix \mathbf{X} , predict the unobserved entries
- Large matrices: # rows, # columns $\approx 10^5, 10^6$.
- Very under-determined (often only 1-2% observed)
- Application to recommender systems, system identification, image processing, microarray data, etc.

Motivations: recommendation systems, Netflix prize

- A popular example is the Netflix challenge (2006-2009)



- Dataset: 480K users, 18K movies, 100M ratings
- Only 1.1% of the matrix is filled!

- In general, we cannot infer missing ratings without any other information.
- This problem is under-determined, more unknown than observations.
- **Low-rank assumption:** fill matrix such that rank is minimum.
→ A few factors explain most of the data.

Completion via rank minimization

$$\text{minimize}_{\mathbf{W}} \text{rank}(\mathbf{W}) \quad \text{s. t.} \quad W_{ij} = \underbrace{X_{ij}}_{\text{observed entries}}, \quad (i, j) \in \underbrace{\Omega}_{\text{sampling set}}.$$

- Non-convex problem and combinatorially NP-hard!!

Convex formulation of the rank minimization problem

$$\text{rank}(\mathbf{X}) = \min_{\mathbf{X}} \sum_{i=1}^{\dim(\mathbf{X})} \mathbb{1}_{(\sigma_i(\mathbf{X}) > 0)} = \|\sigma(\mathbf{X})\|_0.$$

Replace ℓ_0 by ℓ_1 [Fazel (2002), Srebro et al. (2005); Candes and Tao (2010); Recht et al. (2010); Negahban and Wainwright (2011); Klopp (2014)]:

$$\text{Nuclear norm: } \|\mathbf{X}\|_* = \min_{\mathbf{X}} \sum_{i=1}^{\dim(\mathbf{X})} (\sigma_i(\mathbf{X})).$$

Hence tempting to consider

Nuclear norm minimization:

$$\text{minimize}_{\mathbf{W}} \|\mathbf{W}\|_* \quad \text{s. t. } W_{ij} = \underbrace{X_{ij}}_{\text{observed entries}}, \quad (i, j) \in \underbrace{\Omega}_{\text{sampling set}}.$$

This is a convex problem !

Collective matrix completion: motivations

- Data is often obtained from a collection of matrices $\mathcal{X} = (\mathbf{X}^1, \dots, \mathbf{X}^V)$.

$$\mathcal{X} = \left(\begin{array}{ccc} \begin{array}{|c|c|c|c|c|c|} \hline & & \text{green} & & & \\ \hline & \text{orange} & \text{orange} & & & \\ \hline & & \text{yellow} & & & \\ \hline & \text{green} & & & & \\ \hline & & \text{green} & & & \\ \hline & & & & & \\ \hline & & & & & \text{orange} \\ \hline \end{array} & \begin{array}{|c|c|c|c|c|c|} \hline & \text{orange} & & & \text{yellow} & \\ \hline \text{green} & & & & & \\ \hline & & \text{orange} & & & \\ \hline & & & & & \\ \hline & \text{orange} & & & \text{orange} & \\ \hline & & & & \text{yellow} & \\ \hline & & & & & \\ \hline & & & & & \\ \hline \end{array} & \dots & \begin{array}{|c|c|c|c|c|c|} \hline & \text{orange} & & & & \text{orange} \\ \hline & & \text{yellow} & & & \\ \hline & & & & \text{green} & \\ \hline \text{green} & & & & & \\ \hline & \text{orange} & & & & \\ \hline & & & & & \\ \hline & & & & & \text{orange} \\ \hline & & & & & \text{green} \\ \hline \end{array} \\ \mathbf{X}^1 & \mathbf{X}^2 & \dots & \mathbf{X}^V \\ \end{array} \right)$$

- It may be beneficial to leverage all the available user data by various sources.
- Cold-Start** problem: in recommender systems, when a new user has no rating it is impossible to predict his ratings.
- Shared structure among the sources can be useful to get better predictions.

Collective matrix completion: model setup

- Each source view $\mathbf{X}^v \in \mathbb{R}^{d_u \times d_v}$ and $D = \sum_{v=1}^V d_v$.
- We assume that the distribution of for each source \mathbf{X}^v depends on the matrix of parameters \mathbf{M}^v .
- **Model:** let B_{ij}^v be independent Bernoulli random variables and independent from X_{ij}^v , with parameter π_{ij}^v .

$$Y_{ij}^v = B_{ij}^v X_{ij}^v.$$

- We can think of the B_{ij}^v as masked variables.
- $\pi_{ij}^v =$ probability to observe the (i, j) -th entry of the v -th source.

- We consider **general sampling model** where we only assume:

Assumption 1: There exists a positive constant $0 < p < 1$ s.t.
 $\min_{v \in [V]} \min_{(i,j) \in [d_u] \times [d_v]} \pi_{ij}^v \geq p.$

[Klopp (2015); Klopp et al. (2015)]

- $\pi_{i \cdot}^v = \sum_{j=1}^{d_v} \pi_{ij}^v$ the probability to observe an element from the i -th row of \mathbf{X}^v .
- $\pi_{\cdot j}^v = \sum_{i=1}^{d_u} \pi_{ij}^v$ the probability to observe an element from the j -th column of \mathbf{X}^v .

$$\max_{v \in [V]} \max_{(i,j) \in [d_u] \times [d_v]} (\pi_{i \cdot}^v, \pi_{\cdot j}^v) \leq \mu.$$

- Heterogeneous sources: (ratings), (counting: number of clicks) (binomial: like/dislike)
- General framework: natural exponential family:

$$X_{ij}^v | M_{ij}^v \sim h^v(X_{ij}^v) \exp(X_{ij}^v M_{ij}^v - G^v(M_{ij}^v)).$$

[Gunasekar et al. (2014); Cao and Xie (2016); Lafond (2015)]

- Many distributions belong to the exponential family: Gaussian, binomial, Poisson, exponential, etc.

Assumption 2: - The distribution of X_{ij}^v has sub-exponential tail.
- Strong convexity of the log-partition function G^v .

Exponential family noise: estimation procedure

- Given observations $\mathcal{Y} = (\mathbf{Y}^1, \dots, \mathbf{Y}^V)$, we write the negative log-likelihood as

$$\mathcal{L}_{\mathcal{Y}}(\mathcal{W}) = -\frac{1}{d_u D} \sum_{v \in [V]} \sum_{(i,j) \in [d_u] \times [d_v]} B_{ij}^v (Y_{ij}^v W_{ij}^v - G^v(W_{ij}^v)).$$

- The nuclear norm penalized estimator $\widehat{\mathcal{M}}$ of \mathcal{M} is defined as follows:

$$\widehat{\mathcal{M}} = (\widehat{\mathbf{M}}^1, \dots, \widehat{\mathbf{M}}^V) = \underset{\|\mathcal{W}\|_{\infty} \leq \gamma}{\operatorname{argmin}} \mathcal{L}_{\mathcal{Y}}(\mathcal{W}) + \lambda \|\mathcal{W}\|_*,$$

- $\lambda > 0$ is a positive regularization parameter that balances the trade-off between model fit and privileging a low-rank solution.

Upper bound of Frobenius estimation risk norm: the rate of convergence has the following dominant term:

Theorem [A., Klopp 2018]

Assume that Assumptions 1 and 2 hold and

$$\lambda \approx \frac{\sqrt{\mu} + (\log(d_u \vee D))^{3/2}}{d_u D}.$$

Then, with high probability, one has

$$\frac{1}{d_u D} \|\widehat{\mathcal{M}} - \mathcal{M}\|_F^2 \lesssim \frac{\text{rank}(\mathcal{M})(\mu + (\log(d_u \vee D))^{3/2})}{p^2 d_u D}.$$

- **Uniform sampling:** If $c_1/(d_u d_v) \leq \pi_{ij}^v \leq c_2/(d_u d_v)$, then

$$\frac{1}{d_u D} \|\widehat{\mathcal{M}} - \mathcal{M}\|_F^2 \lesssim \frac{\text{rank}(\mathcal{M})}{p(d_u \wedge D)}.$$

- We denote $n = \sum_{v \in [V]} \sum_{(i,j) \in [d_u] \times [d_v]} \pi_{ij}^v$, the expected number of observations.
- **Sample complexity:**

$$n \gtrsim \text{rank}(\mathcal{M})(d_u \vee D).$$

Example: 1-bit matrix completion

- **1-bit matrix completion:** $\mathcal{Y} \in \{+1, -1\}$ with probability $f(\mathcal{M})$ for some link-function f [Davenport et al. (2014); Klopp et al. (2015); Alquier et al. (2017)]
- Klopp et al. (2015) obtained the rate $\text{rank}(\mathcal{M})(d_u \vee D) \log(d_u \vee D)/n$ as the upper bound and $\text{rank}(\mathcal{M})(d_u \vee D)/n$ as the lower bound for 1-bit matrix completion.

Corollary[A., Klopp 2018]

$$\frac{1}{d_u D} \|\widehat{\mathcal{M}} - \mathcal{M}\|_F^2 \lesssim \frac{\text{rank}(\mathcal{M})(d_u \vee D)}{n},$$

- **Answer** the important theoretical question: what is the exact minimax rate of convergence for 1-bit matrix completion which was previously known up to a logarithmic factor.

- We do not assume any specific model for the observations.
- We consider the risk of estimating \mathbf{X}^ν with a loss function ℓ^ν ,
- We focus on non-negative loss functions that are Lipschitz:

Assumption 3: We assume that the loss function $\ell^\nu(y, \cdot)$ is ρ_ν -Lipschitz in its second argument:

$$|\ell^\nu(y, x) - \ell^\nu(y, x')| \leq \rho_\nu |x - x'|.$$

- Examples: hinge loss with $\ell^\nu(y, y') = \max(0, 1 - yy')$, logistic loss with $\ell^\nu(y, y') = \log(1 + \exp(-yy'))$, etc.

- Goodness-of-fit term:

$$R_{\mathbf{y}}(\mathcal{W}) = \frac{1}{d_u D} \sum_{v \in [V]} \sum_{(i,j) \in [d_u] \times [d_v]} B_{ij}^v \ell^v(Y_{ij}^v, W_{ij}^v).$$

- We define the oracle as:

$$\hat{\mathcal{M}}^* = (\hat{\mathcal{M}}^{*1}, \dots, \hat{\mathcal{M}}^{*V}) = \underset{\|\mathcal{W}\|_\infty \leq \gamma}{\operatorname{argmin}} R(\mathcal{W}),$$

where $R(\mathcal{W}) = \mathbb{E}[R_{\mathbf{y}}(\mathcal{W})]$.

- For a tuning parameter $\Lambda > 0$, the nuclear norm penalized estimator $\hat{\mathcal{M}}$ is defined as

$$\hat{\mathcal{M}} \in \underset{\|\mathcal{W}\|_\infty \leq \gamma}{\operatorname{argmin}} \{R_{\mathbf{y}}(\mathcal{W}) + \Lambda \|\mathcal{W}\|_*\}.$$

- We denote by $\|\mathcal{W}\|_{\Pi, F}^2 = \sum_v \sum_{(i,j)} \pi_{ij}^v (W_{ij}^v)^2$.

Assumption 4: Assume that for every \mathcal{W} with $\|\mathcal{W}\|_{\infty} \leq \gamma$, one has $R(\mathcal{W}) - R(\mathcal{M}^{\star}) \gtrsim \frac{1}{d_u D} \|\mathcal{W} - \mathcal{M}^{\star}\|_{\Pi, F}^2$.

- Assumption 4 is called “Bernstein” condition (Mendelson, 2008; Bartlett et al., 2004; Alquier et al., 2017; Elsen and van de Geer, 2018).

Theorem [A. Klopp 2018]

Let Assumptions 1, 3, and 4 hold and $\Lambda \approx (\sqrt{\mu} + \sqrt{\log(d_u \vee D)}) / (d_u D)$. Then, with probability , one has

$$R(\widehat{\mathcal{M}}) - R(\mathcal{M}^{\star}) \lesssim \frac{\mu + \log(d_u \vee D)}{pd_u D}$$

- First theoretical guarantees on the case of noisy collective MC.
- Collective approach provides faster rate of convergences in the case of joint low-rank structure.
- Exact minimax optimal rate of convergence for 1-bit matrix completion which was known upto a logarithmic factor.
- On going work: algorithmic study with numerical experiments.



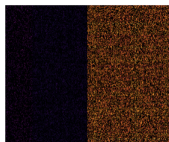
Thank You.

Toy Illustration using cvxpy package

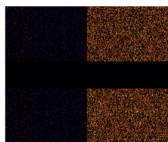
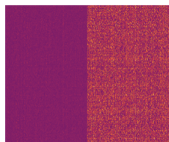
V	d_u	d_1	d_2	d_3	$M^1(\text{rank} = 5)$	$M^2(\text{rank} = 10)$	$M^3(\text{rank} = 15)$
3	500	100	200	300	$\mathcal{N}(-2, 0.5)$	$\mathcal{N}(1, 0.5)$	$\mathcal{N}(2, 0.5)$

	M^1	M^2	M^3	\mathcal{M}	$\mathcal{M}_{\text{cold}}$
% observations	10%	20%	30%	23.29%	18.69%

	CMC	SNN	Cold-Start	\widehat{M}^1	\widehat{M}^2	\widehat{M}^3
RMSE	0.223	0.224	0.220	0.198	0.194	0.311



observed + fitted collective matrix.



observed + fitted cold collective matrix.

cvxpy [Diamond and S. Boyd (2016)]

Alquier, P., V. Cottet, and G. Lecué (2017). Estimation bounds and sharp oracle inequalities of regularized procedures with Lipschitz loss functions. *arXiv:1702.01402*.

Bartlett, P. L., M. I. Jordan, and J. D. McAuliffe (2004). Large margin classifiers: Convex loss, low noise, and convergence rates. In S. Thrun, L. K. Saul, and B. Schölkopf (Eds.), *Advances in Neural Information Processing Systems 16*, pp. 1173–1180. MIT Press.

Candes, E. J. and T. Tao (2010). The power of convex relaxation: Near-optimal matrix completion. *IEEE Transactions on Information Theory* 56(5), 2053–2080.

Cao, Y. and Y. Xie (2016, March). Poisson matrix recovery and completion. *IEEE Transactions on Signal Processing* 64(6), 1609–1620.

Davenport, M. A., Y. Plan, E. van den Berg, and M. Wootters (2014). 1-bit matrix completion. *Information and Inference: A Journal of the IMA* 3(3), 189.

Elsener, A. and S. van de Geer (2018). Robust low-rank matrix

estimation. *To appear in The Annals of Statistics, arXiv preprint arXiv:1603.09071.*

Fazel, M. (2002). *Matrix Rank Minimization with Applications*. Ph. D. thesis, Stanford University.

Gunasekar, S., P. Ravikumar, and J. Ghosh (2014). Exponential family matrix completion under structural constraints. In *Proceedings of the 31st International Conference on International Conference on Machine Learning - Volume 32, ICML'14*, pp. II–1917–II–1925. JMLR.org.

Klopp, O. (2014). Noisy low-rank matrix completion with general sampling distribution. *Bernoulli* 20(1), 282–303.

Klopp, O. (2015). Matrix completion by singular value thresholding: Sharp bounds. *Electron. J. Statist.* 9(2), 2348–2369.

Klopp, O., J. Lafond, E. Moulines, and J. Salmon (2015). Adaptive multinomial matrix completion. *Electron. J. Statist.* 9(2), 2950–2975.

Lafond, J. (2015, 03–06 Jul). Low rank matrix completion with exponential family noise. In P. Grünwald, E. Hazan, and S. Kale

(Eds.), *Proceedings of The 28th Conference on Learning Theory*, Volume 40 of *Proceedings of Machine Learning Research*, Paris, France, pp. 1224–1243. PMLR.

Mendelson, S. (2008). Obtaining fast error rates in nonconvex situations. *Journal of Complexity* 24(3), 380 – 397.

Negahban, S. and M. J. Wainwright (2011). Estimation of (near) low-rank matrices with noise and high-dimensional scaling. *Ann. Statist.* 39(2), 1069–1097.

Recht, B., M. Fazel, and P. A. Parrilo (2010). Guaranteed minimum-rank solutions of linear matrix equations via nuclear norm minimization. *SIAM Rev.* 52(3), 471–501.

Srebro, N., J. Rennie, and T. S. Jaakkola (2005). Maximum-margin matrix factorization.